

Successive events modeling

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Introduction

Failure

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Disaster

Introduction

Failure

random occurrence in time
independent recurrence
renewal process
quick repair or renewal

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Disaster

random occurrence in time
growing spread
domino effect
repair slower than spreading

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equipment failure
 accidents, health damage
 information system failure
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 natural accident

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Model → Prediction → Prevention

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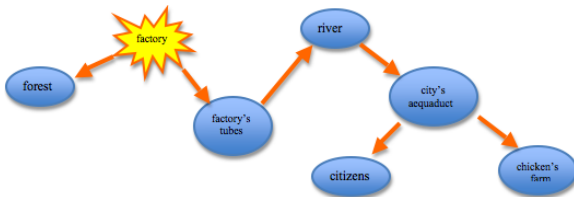
Model → Prediction → Prevention

Preventive Maintenance Policy

Disaster Recovery Plan

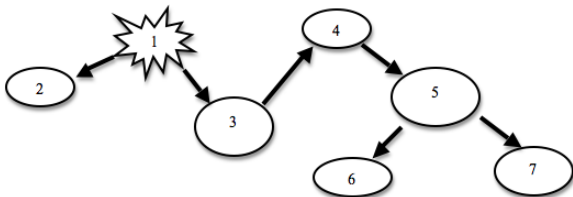
Introduction

The disastrous model (ESREL, September 2008):



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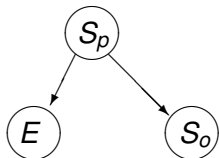
The disastrous model (ESREL, September 2008):



State of the system: $\omega(t) = (\omega_1(t), \omega_2(t), \dots, \omega_n(t)), \quad t \geq 0,$

- (i) at the beginning (at the time 0), the system is in the state $\omega^0 = (0, 0, \dots, 0),$
- (ii) the process starts by deterioration of an object i with probability $\pi_i, i = 1, 2, \dots, n,$
- (iii) when at the time t an object i was affected, there was a random time period τ after which the event moved onto some of the unaffected objects,
- (iv) states of the system in time t create a stochastic process $\{X(t), t \geq 0\}$ in continuous time. Values of this process lie within the set $\Omega = \{0, 1\}^n$

SET model

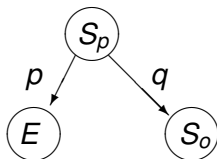


Successive event chain (SEC)

SET model

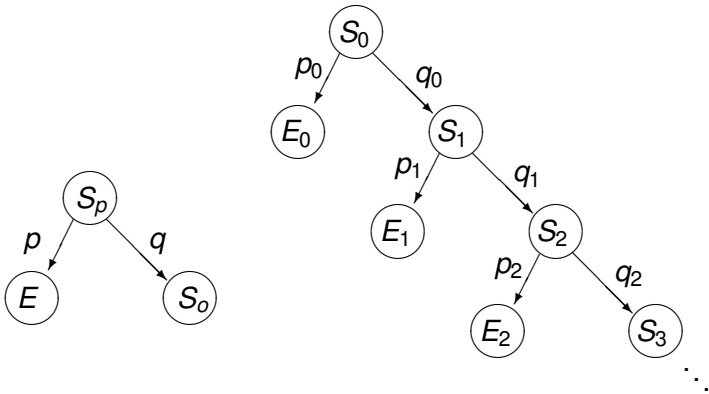
At any time $t > 0$, SEC will make decisions on the following three possibilities:

- to leave the state S_p and continue to a offspring event S_o with intensity q
- to finish the sequence with intensity p .
- to stay in the state S_p for a random dwell time T , which has an exponential distribution $\text{Exp}(-(q + p))$



Successive event chain (SEC)

SET model



Successive event chain (SEC) – successive event tree (SET)

SET model

$\mathcal{S} = \{S_0, S_1, \dots, S_n\}$ as the set of successive events,
 $\mathcal{E} = \{E_0, E_2, \dots, E_n\}$ as the set of termination events.

We shall describe the SET as a homogenous Markov process $\{X(t), t \geq 0\}$ with continuous time and a finite set of states $\mathcal{S} \cup \mathcal{E}$.

The process starts in the state S_0 with probability 1.
All states in \mathcal{S} are transient and the states in \mathcal{E} are trapping.

$q_{i,i+1} = q_i$ transition intensity from state S_i to S_{i+1} ,
 $q_{i,n+i+2} = p_i$ transition intensity from S_i to absorbing state E_i .

SET model

At any time $t > 0$, SEC will make decisions on the following three possibilities:

- to stay in the state S_i for a random dwell time T_i , which has an exponential distribution $\text{Exp}(-(q_i + p_i))$
- to leave the state S_i and continue to a offspring event S_{i+1} with intensity q_i
- to finish the sequence with intensity p_i in a state E_i .

SET model

The transition intensity matrix has the block form

$$Q = \begin{pmatrix} Q_S & Q_E \\ O & O \end{pmatrix} \quad (1)$$

where Q_E is a $(n+1) \times (n+1)$ diagonal matrix whose elements p_i are the intensities of stopping after the i th successive event S_i , $i = 0, \dots, n$ and O denotes $(n+1) \times (n+1)$ null matrix. Q_S is the $(n+1) \times (n+1)$ transition intensity matrix from states in S to states in S in the form

$$Q_S = \begin{pmatrix} -(q_0 + p_0) & q_0 & 0 & \cdots & 0 \\ 0 & -(q_1 + p_1) & q_1 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & -p_n \end{pmatrix}$$

SET model

Let us consider the embedded Markov chain $\{Y_n\}$, $n = 0, 1, \dots, n$. The transition probability matrix of Y in the block form is

$$P = \begin{pmatrix} P_S & P_E \\ O & I \end{pmatrix}$$

where

$$P_S = \begin{pmatrix} 0 & \frac{q_0}{(q_0+p_0)} & 0 & \cdots & 0 \\ 0 & 0 & \frac{q_1}{(q_1+p_1)} & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & 0 \end{pmatrix},$$

where P_E is a $(n+1) \times (n+1)$ diagonal matrix whose elements $p_i/(q_i + p_i)$ are the probabilities of stopping after the i -th successive event S_i , $i = 0, \dots, n$. O denotes a $(n+1) \times (n+1)$ null matrix and I is the $(n+1) \times (n+1)$ identity matrix.

Proposition

Let $\{X(t), t \geq 0\}$ is a SET with transition intensity matrix Q .
Then $X(t)$ stops in the state k with probabilities

$$\begin{aligned}\pi_0 &= \frac{p_0}{p_0 + q_0} = 1 - \frac{q_0}{p_0 + q_0}, \\ \pi_k &= p_k \cdot \frac{\prod_{j=0}^{k-1} q_j}{\prod_{j=0}^k (p_j + q_j)}, \quad k = 1, \dots, n-1, \\ \pi_n &= \frac{\prod_{j=0}^{n-1} q_j}{\prod_{j=0}^{n-1} (p_j + q_j)}.\end{aligned}$$

The mean time to stopping after the i th state is

$$\tau_k = \sum_{j=0}^k \frac{1}{p_j + q_j}.$$

Example

Let us consider a SET which is assembled from n identical SECs. This means that we have $q_i = q$, $p_i = p$ for all $i = 0, \dots, n$. Then the probability distribution of terminal states \mathcal{E} has the form

$$\pi_k = \left(\frac{q}{p+q}\right)^{k-1} \left(1 - \frac{q}{p+q}\right), \quad 1 \leq k \leq n-1,$$
$$\pi_n = \left(\frac{q}{p+q}\right)^{n-1}$$

and the mean stop-time is $T_k = \frac{k}{p+q}$. The mean "duration" of such finite SET is $\frac{1}{p} \left(1 - \frac{q^n}{(p+q)^n}\right) = \frac{1}{p} (1 - \pi_n)$.

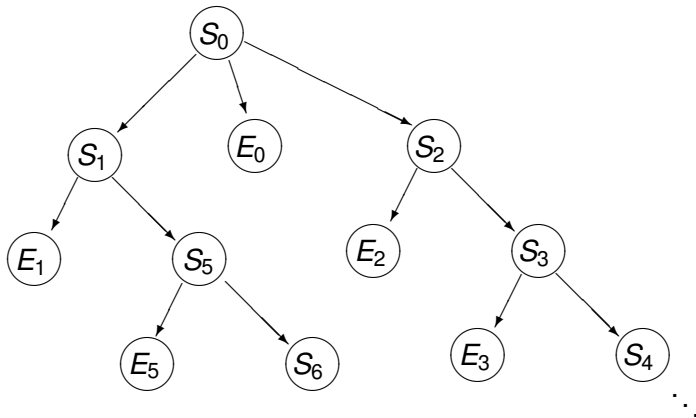
A note: Semi-markovian model

Consider a model in which dwelt times are not necessary exponential. It means that we suppose a set of given distributive functions $F_i(t)$, $t \geq 0$ of dwelt times in states S_i , $i = 1, \dots, n$, and the process $X(t)$ to be semi-markovian. When the embedded Markov chain remains the same, the previous Proposition holds but the mean time to stopping will be

$$T_k = \sum_{j=0}^k E(\tau_j),$$

where $E(\tau_j) = \int_0^{\infty} t dF_j(t)$.

One-way branching model



Proposition

Let $\{X(t), t \geq 0\}$ be a SET with $(n \times n)$ transition intensity matrix Q in the previous form where Q_S and Q_E are upper triangular matrices. Denote P the transition probability matrix of an embedded Markov chain.

Then $X(t)$ stops in the state k with probabilities

$$\vec{\pi} = \vec{a}P_S^*P_E,$$

where $\vec{a} = (1, 0, \dots, 0)$ is a initial probability vector of $X(t)$ and

$$P_S^* = \sum_{j=1}^n P_S^j.$$

The mean time to stopping in state E_k is given by the formula

$$T_k = -\vec{a}(Q_S^{(k)})^{-1}\vec{e}',$$

where $\vec{e} = (1, 1, \dots, 1)$ and $Q_S^{(k)}$ equals to the matrix Q_S but diagonal elements are

$$q_{i,i}^{(k)} = q_{i,i} + \sum_{j \neq k} q_{i,j}.$$

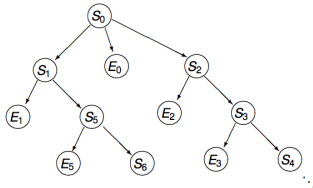
Semi-markovian model

In a semi-markov case, the situation is more complicated.

Suppose that we are able to approximate distribution functions F_j by phase-type distributions with representations (B_j, b_j^0) , $j = 1, \dots, n$ (see Neuts, 1975). Then the transition intensity matrix has the form

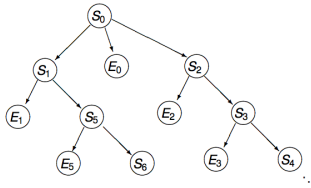
Conclusions

- There has been proposed the model for successive events spreading in two variants:



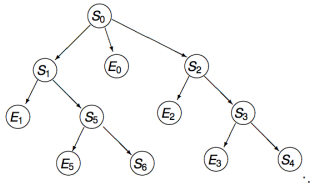
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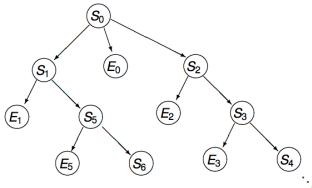
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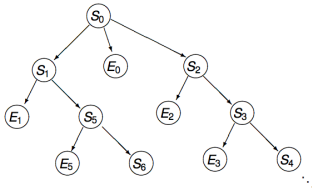
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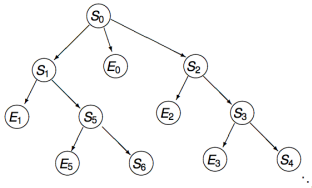
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- **The probability distribution of the length of spreading is provided.**



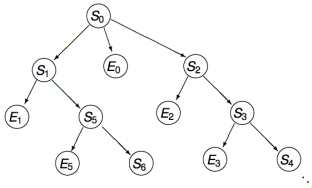
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Thanks for your attention.